

# Week 3: Simple Linear Regression

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# Outline

- Putting a structure into exploratory data analysis
- Covariance and correlation
- Simple linear regression
- Parameter estimation
- Regression towards the mean

# Big picture

- We did exploratory data analysis of two variables  $X$ ,  $Y$  in the first homework
- Now we are going to provide **structure** to the analysis
- We will assume a relationship (i.e. a **functional form**) and estimate parameters that **summarize** that relationship
- We will then **test** hypotheses about the relationship
- For the time being, we will focus on **two continuous** variables

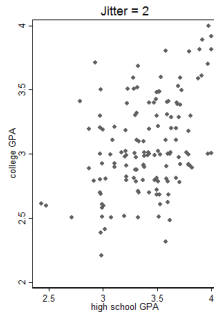
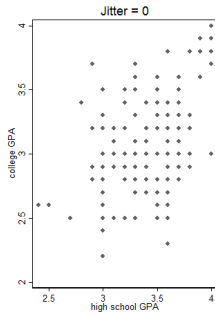
# Example data

- We will use data from Wooldridge on grades for a sample 141 college students (see today's do file)

```
use GPA1.DTA, clear
```

```
sum colgpa hsgpa
```

Variable	Obs	Mean	Std. Dev.	Min	Max
colgpa	141	3.056738	.3723103	2.2	4
hsgpa	141	3.402128	.3199259	2.4	4

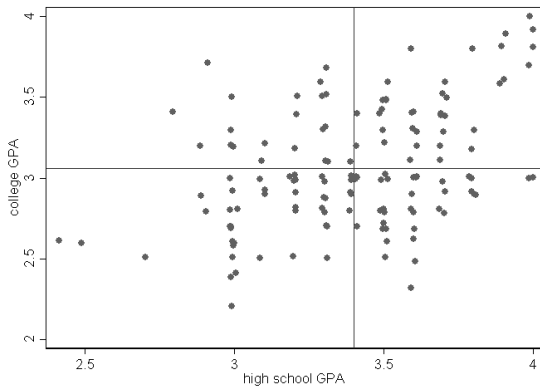


## Covariance and correlation

- A simple summary of the relationship between two variables is the **covariance**:
- $COV(X, Y) = E[(X - E(X))(Y - E(Y))] = E(XY) - E(X)E(Y)$
- $COV(X, Y) = \frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y})(x_i - \bar{x})$
- For each pair  $x_i, y_i$  we calculate the product of the deviations of each variable from its mean
- The covariance will be closer to zero if observations are closer to their mean (for one or both variables); it can be positive or negative
- The scale is the product of the scales of  $X$  and  $Y$  (e.g. age\* age, grades\*age, etc)

```
. corr colgpa hsgpa, c
(obs=141)
-----+-----
      |      colgpa      hsgpa
colgpa |      .138615
hsgpa  |      .049378      .102353
```

## Graphical intuition?



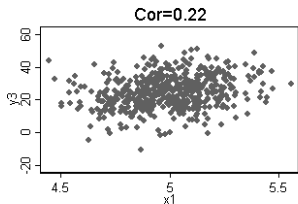
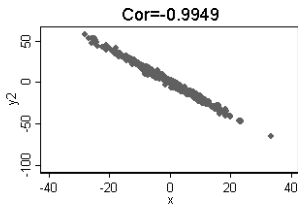
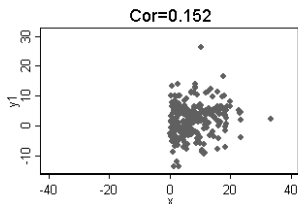
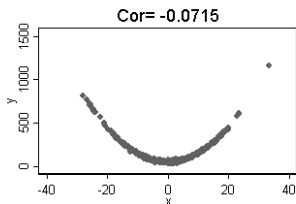
- If  $COV(X, Y) > 0$  then positive relationship between  $x$  and  $y$
- If  $COV(X, Y) < 0$  then negative relationship between  $x$  and  $y$

# Correlation

- The sign of the covariance is useful but the magnitude is not because it **depends on the unit of measurement**
- The **correlation** ( $\rho$ ) scales the covariance by the standard deviation of each variable:
- $$Cor(X, Y) = \frac{1}{n-1} \sum_{i=1}^n \frac{(y_i - \bar{y})(x_i - \bar{x})}{S_y S_x}$$
- $$Cor(X, Y) = \frac{COV(X, Y)}{S_y S_x}$$
- $-1 \leq Cor(X, Y) \leq 1$
- Closer to 1 or -1, stronger relationship
- Grades data:

corr	colgpa	hsgpa	
			colgpa hsgpa
-----			
	colgpa		1.0000
	hsgpa		0.4146 1.0000

# Examples



- $\rho$  close to 0 **does NOT imply**  $X$  and  $Y$  are not related
- $\rho$  measures the **linear** relationship between two variables



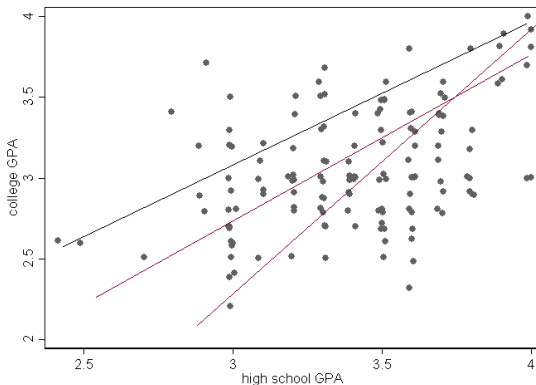
## Going beyond the correlation coefficient

- We need more flexibility to understand the relationship between  $X$  and  $Y$ ; the correlation is useful but it is limited to a linear relationship and we can't study changes in  $Y$  for changes in  $X$  using  $\rho$
- A useful place to start is assuming a more specific functional form:
- $Y = \beta_0 + \beta_1 X + \epsilon$
- The above model is the an example of **simple linear regression** (SLR)
- **Confusion alert:** it's **linear on the parameters**  $\beta_i$ ;  
 $Y = \beta_0 + \beta_1 X^2 + \epsilon$  is also a SLR model
- In the college grades example, we have  $n = 141$  observations. We could write the model as
- $y_i = \beta_0 + \beta_1 x_i + \epsilon_i$ , where  $i = 1, \dots, n$ . College grades is  $y$  and high school grades is  $x$

## The role of $\epsilon$

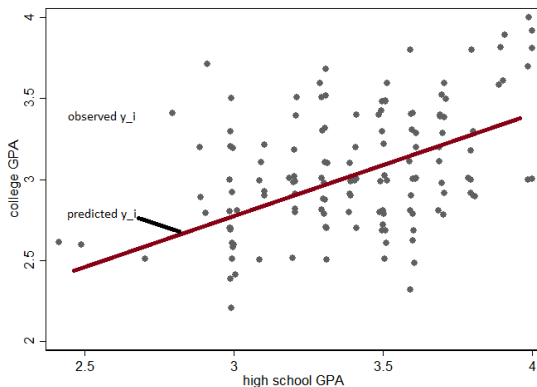
- $y_i = \beta_0 + \beta_1 x_i + \epsilon_i$ , where  $i = 1, \dots, n$
- We are assuming that  $X$  and  $Y$  are related as described by the above equation plus an error term  $\epsilon$
- In general, we want the error, or the **unexplained part of the model**, to be as small as possible
- How do we find the optimal  $\beta_j$ ? One way is to find the values of  $\beta_0$  and  $\beta_1$  that are **as close as possible** to **all** the points  $x_i, y_i$
- These values are  $\hat{\beta}_0$  and  $\hat{\beta}_1$  the prediction is  $\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x$
- This is equivalent to say that we want to make the  $\epsilon$  as small as possible
- Obviously, the relationship is not going to be perfect so  $\epsilon_i \neq 0$  for most observations

## Some possible lines (guesses)



- I used a graphic editor to draw some possible lines; I wanted to draw the lines as close as possible to most of the points
- The line is affected by the mass of points and extreme values

## A more systematic way



- The error will be the difference  $\epsilon = (y_i - \hat{y}_i)$  for each point; we don't want a positive error to cancel out a negative one so we take the square:  $\epsilon_i^2 = (y_i - \hat{y}_i)^2$

## The method of ordinary least squares (OLS)

- We want to find  $\hat{\beta}_i$  that minimizes the sum of all errors:

$$S(\beta_0, \beta_1) = \sum_{i=1}^n \epsilon_i^2 = \sum_{i=1}^n (y_i - \hat{y}_i)^2 = \sum_{i=1}^n (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2$$

- The solution is fairly simply with calculus. We solve the system of equations:

$$\frac{\partial S(\beta_0, \beta_1)}{\partial \beta_0} = 0$$

$$\frac{\partial S(\beta_0, \beta_1)}{\partial \beta_1} = 0$$

- The solution is  $\hat{\beta}_1 = \frac{\sum_{i=1}^n (y_i - \bar{y})(x_i - \bar{x})}{\sum_{i=1}^n (x_i - \bar{x})^2}$  and  $\beta_0 = \bar{y} - \hat{\beta}_1 \bar{x}$
- To get predicted values, we use  $\hat{y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_i$
- $S(\beta_0, \beta_1)$  is also denoted by SSE, *sum of squares for error*

## Deriving the formulas

- We start with:

$$\frac{\partial SSE}{\partial \beta_0} = \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i)^2 = -2 \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i) = 0$$

- We can then multiply by  $-\frac{1}{2}$  and distribute the summation:

$$\sum_{i=1}^n y_i - n\beta_0 - \beta_1 \sum_{i=1}^n x_i = 0$$

- And almost there. Divide by  $n$  and solve for  $\beta_0$ :  $\hat{\beta}_0 = \bar{y} - \beta_1 \bar{x}$

- For  $\beta_1$ , more steps but start with the other first order condition and plug in  $\hat{\beta}_0$

$$\frac{\partial SSE}{\partial \beta_1} = -2 \sum_{i=1}^n x_i (y_i - \beta_0 - \beta_1 x_i) = 0$$

## Interpreting the formulas

- Does the formula for  $\hat{\beta}_1$  look familiar?

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (y_i - \bar{y})(x_i - \bar{x})}{\sum_{i=1}^n (x_i - \bar{x})^2}$$

- We can multiply by  $\frac{1/(n-1)}{1/(n-1)}$  and we get the formulas for the covariance and variance:

$$\hat{\beta}_1 = \frac{COV(Y, X)}{Var(X)}$$

- Since  $Var(X) > 0$ , the sign of  $\hat{\beta}_1$  depends on  $COV(X, Y)$
- If the X and Y are not correlated, then  $\hat{\beta}_1 = 0$
- So you can use a test for  $\hat{\beta}_1 = 0$  as a **test for correlation**. But now you have more flexibility and are not constrained to a linear relationship correlation
- For example, you could test if  $\gamma_1 = 0$  in  $Y = \gamma_0 + \gamma_1 X^2$

## Digression

- Not the only target function to minimize. We could also work with the absolute value, as in  $|y_i - \hat{y}_i|$ . This is called the least absolute errors regression; more robust to extreme values
- **Jargon alert:** *robust* means a lot of things in statistics. Whenever you hear that XYZ method is more robust, ask the following question: robust to what? It could be missing values, correlated errors, functional form...
- A very fashionable type of model in prediction and machine learning is the **ridge regression** (Lasso method, too)
- It minimizes the sum of errors  $(y_i - \hat{y}_i)^2$  plus the sum of square betas  $\lambda \sum_{j=1}^j \beta_j^2$
- It may look odd but we want to also make the betas as small as possible as a way to **select variables** in the model



## Grades example

- In Stata, we use the reg command:

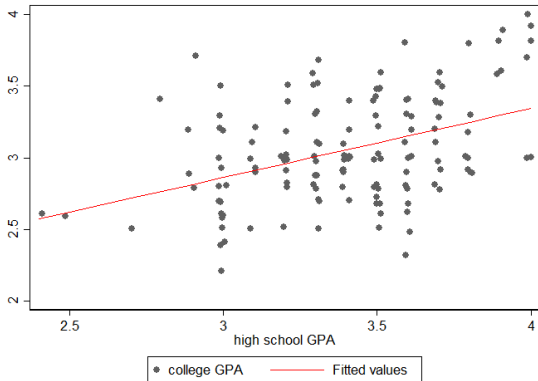
```
. reg colgpa hsgpa
```

Source	SS	df	MS	Number of obs	=	141
-----+				F(1, 139)	=	28.85
Model	3.33506006	1	3.33506006	Prob > F	=	0.0000
Residual	16.0710394	139	.115618988	R-squared	=	0.1719
-----+				Adj R-squared	=	0.1659
Total	19.4060994	140	.138614996	Root MSE	=	.34003
-----						
colgpa	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+						
hsgpa	.4824346	.0898258	5.37	0.000	.304833	.6600362
_cons	1.415434	.3069376	4.61	0.000	.8085635	2.022304
-----						

- So  $\hat{\beta}_0 = 1.415434$  and  $\hat{\beta}_1 = .4824346$ . A predicted value is  $\hat{y}_i = 1.415434 + .4824346(x_i = a)$

## Grades example II

```
gen gpahat = 1.415434 + .4824346*hsgpa
gen gpahat0 = _b[_cons] + _b[hsgpa]*hsgpa
predict gpahat1
* ereturn list
* help reg
scatter colgpa hsgpa, jitter(2) || line gpahat1 hsgpa, color(red) sort ///
    saving(reg1.gph, replace)
```

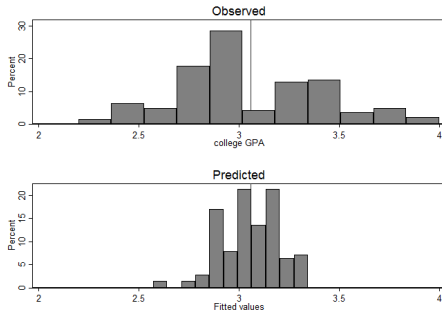


# How do observed and predicted values compare?

```
sum colgpa gpahat1
hist colgpa, percent title("Observed") saving(hisob.gph, replace) xline(3.06)
hist gpahat1, percent title("Predicted") saving(hispred.gph, replace) xline(3.06)
graph combine hisob.gph hispred.gph, col(1) xcommon
```

Variable	Obs	Mean	Std. Dev.	Min	Max
colgpa	141	3.056738	.3723103	2.2	4
gpahat1	141	3.056738	.1543433	2.573277	3.345172

## ■ Predictions “regress” towards the mean:



## Regression towards the mean

- Regression towards the mean is an often-misunderstood concept
- In this example, our model is telling us that a student with a high high-school GPA is going to be more like an average college student (i.e. she will regress towards the mean)
- Why is that happening? Look at the data. Is that true in our sample?
- It happens because our prediction is using the information of **everybody** in the sample to make predictions for those with high high-school GPA
- It may also be because it's a *property* of the particular dataset or problem, like in the homework example

## Confusion alert and iterated expectations

- From OLS, it is not clear that we are modeling the conditional expectation of  $Y$  given  $X$ :  $E[Y|X]$  but **WE ARE (!!)**
- We are modeling how the mean of  $Y$  changes for different values of  $X$
- The **mean of the predictions** from our model **will match the observed mean of  $Y$**
- We can use the law of iterated expectations to go from the conditional to unconditional mean of  $Y$ :

```
qui sum hsgpa
di 1.415434 + .4824346*r(mean)
3.0567381
```

```
.sum colgpa
```

Variable	Obs	Mean	Std. Dev.	Min	Max
colgpa	141	3.056738	.3723103	2.2	4

## Another way of writing the model

- When we cover Maximum Likelihood Estimation (MLE), it's going to become super clear that we are indeed modeling a **conditional expectation**
- For the rest of the semester and your career, it would be useful to write the estimated model as  $E(\hat{y}_i|x) = \hat{\beta}_0 + \hat{\beta}_1x$  or  $E(\hat{y}_i) = \hat{\beta}_0 + \hat{\beta}_1x$
- Next class we are going to start interpreting parameters. We will see that  $\hat{\beta}_1$  tells you how the expected value/average  $y$  changes when  $x$  changes
- This is **subtle** but super important. It's not the change in  $y$ , it's the change in the average  $y$
- Seeing it this way will make it easier later (trust me)
- To make it a bit more confusing: of course, we can use the model to make a prediction for one person. Say, a college student with a hs gpa of  $xx$  will have a college gpa of  $yy$ . But that prediction is based on the average of others

# Big picture

- We started with a graphical approach to study the relationship of two continuous variables
- We then used the correlation coefficient to measure the **magnitude** and **direction** of the **linear** relationship
- We then considered a more **flexible** approach by assuming a more specific functional form and used the method of least squares to find the best parameters
- We now have a way of **summarizing** the relationship between  $X$ ,  $Y$
- We **didn't make any assumptions about the distribution** of  $Y$  (or  $X$ )
- Don't ever forget that we are modeling the conditional expectation (!!)
- Next class we will see other ways of thinking about SLR and causal inference